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## ASSUMPTION-LEAN QUANTILE REGRESSION

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# THE MODELING TRADITION



## THE MODELING TRADITION

## Statistical modeling

#### Models are central

#### Recent critiques:

(Breiman, 2001; Freedman, 2001; Robins & Rotnitzky, 2001; van der Laan, 2015; ...)

- Occam's dilemma: simple & interpretable vs complex & plausible.
- We need to make compromises
  - $\Rightarrow$  misspecification and bias.
- Model building ⇒ bias and post-selection

inference (Leeb & Pötscher, 2006; Dukes & Vansteelandt, 2020)

#### Algorithmic modeling

- Model misspecification is much less a concern.
- But focus is on prediction.
- Not aimed at explanation.
- No real uncertainty assessments.

The causal modeling culture increasingly builds on the algorithmic culture, instead targeting model-free estimands and providing valid uncertainty assessments.

# How can we bridge These modeling cultures?



## ASSUMPTION-LEAN REGRESSION (1)

That is what is achieved in a recent JRSS B discussion paper on assumption-lean modeling.

Vansteelandt S, Dukes O. Assumption-lean inference for generalised linear model parameters (with discussion). JRSS-B 2022.

Consider the semi-parametric structural quantile model

$$\underbrace{\mathcal{Q}_{\tau}(Y^{a}|L)}_{\mathcal{Q}_{\tau}(Y|A=a,L)} - \underbrace{\mathcal{Q}_{\tau}(Y^{0}|L)}_{\text{unknown fct of }L} = \beta_{\tau}(L)a \quad \text{for all } a.$$

- Assume that adjustment for *L* suffices to control for confounding:  $Y^a \perp \!\!\!\perp A | L$ .
- Techniques for partially linear quantile models are relevant, but have limited utility: (Lee, 2003; Sun, 2005; Wu et al., 2010; Wu and Yu, 2014; Lv et al., 2015; Sherwood and Wang, 2016; Zhong and Wang, 2023)
  - computational demands;
  - challenges in high-dimensional applications (due to reliance on kernel weighting or splines);
  - biased inference when the model is wrong.

## ASSUMPTION-LEAN REGRESSION (2)

Because the model

$$Q_{ au}(Y^a|L) - Q_{ au}(Y^0|L) = eta_{ au}(L)a$$
 for all  $a$ 

is deliberately kept simple, we will not assume it to hold, but use it to communicate our results.

- The real modeling is done through statistical / machine learning, results of which are projected and de-biased in view of a specific estimand.
- As such, we ensure that we are estimating a well-understood exposure effect and obtain valid inferences, even when the model is misspecified, and despite the use of machine learning.

# ASSUMPTION-LEAN QUANTILE REGRESSION



## BE CLEAR ABOUT THE ESTIMAND (1)

- A 'hygienic' analysis is clear about the estimand, even when models are used.
- For instance, with a binary randomized treatment A, we map  $\beta_{\tau}(L)$  in model

$$Q_{\tau}(Y^{1}|L) - Q_{\tau}(Y^{0}|L) = \beta_{\tau}(L)$$

onto the model-free estimand

$$\mathbb{E}\left\{ Q_{ au}(Y^{1}|L) - Q_{ au}(Y^{0}|L) 
ight\},$$

which is what we will estimate.

- This choice prevents that naïve interpretation as a 'difference between quantiles' would be misleading.
- In contrast, in standard (partially linear) quantile regression, it is unclear what we are estimating when the model is wrong.

When *A* is not randomized, we may consider the same estimand, or generalize it to the weighted average:

$$\frac{\mathbb{E}[w(L)\left\{Q_{\tau}(Y^{1}|L)-Q_{\tau}(Y^{0}|L)\right\}]}{\mathbb{E}\left\{w(L)\right\}},$$

with

$$w(L) = P(A = 1|L)P(A = 0|L).$$

## DEBIASED MACHINE LEARNING



### A DEBIASED ESTIMATOR

• When  $Y^a \perp A \mid L$ , the estimand can be identified as

$$\frac{\mathbb{E}\left(\left\{A - \mathbb{E}(A|L)\right\}\left[Q_{\tau}(Y|A,L) - \mathbb{E}\left\{Q_{\tau}(Y|A,L)|L\right\}\right]\right)}{\mathbb{E}\left[\left\{A - \mathbb{E}(A|L)\right\}^{2}\right]}$$

Based on the estimand's efficient influence function, we construct the following debiased estimator

$$\frac{1}{n} \sum_{i=1}^{n} \frac{A_{i} - \hat{\mathbb{E}}(A_{i}|L_{i})}{\frac{1}{n} \sum_{i=1}^{n} \{A_{i} - \hat{\mathbb{E}}(A_{i}|L_{i})\}^{2}} \left[ \hat{Q}_{\tau}(Y_{i}|A_{i}, L_{i}) - \hat{\mathbb{E}} \Big\{ \hat{Q}_{\tau}(Y_{i}|A_{i}, L_{i}) \Big| L_{i} \Big\} \right] \\ + \frac{1}{n} \sum_{i=1}^{n} \frac{A_{i} - \hat{\mathbb{E}}(A_{i}|L_{i})}{\frac{1}{n} \sum_{i=1}^{n} \{A_{i} - \hat{\mathbb{E}}(A_{i}|L_{i})\}^{2}} \left[ \frac{\tau - I\{Y_{i} \leq \hat{Q}_{\tau}(Y_{i}|A_{i}, L_{i})\}}{\hat{f}_{Y|A,L}(\hat{Q}_{\tau}(Y_{i}|A_{i}, L_{i})|A_{i}, L_{i})} \right],$$

where the nuisance parameters are substituted by data-adaptive estimates (e.g., ML).

## A TARGETED LEARNING ESTIMATOR

- Targeted learning 'simplifies' this by forcing the second line to give zero, which gives an asymptotically equivalent estimator.
- It does so by 'targeting' an initial estimator  $\widetilde{Q}_{\tau}(Y|A, L)$  so that

$$\frac{1}{n}\sum_{i=1}^{n}\left\{A_{i}-\hat{\mathbb{E}}(A_{i}|L_{i})\right\}\left[\frac{\tau-l\{Y_{i}\leq\widetilde{Q}_{\tau}(Y_{i}|A_{i},L_{i})\}}{\hat{\mathfrak{f}}_{Y|A,L}(\widetilde{Q}_{\tau}(Y_{i}|A_{i},L_{i})|A_{i},L_{i})}\right]\approx0.$$

This is done by fitting the quantile regression model

$$\widetilde{Q}_{\tau}(\mathbf{Y}_{i}|\mathbf{A}_{i}, L_{i}) = \hat{Q}_{\tau}(\mathbf{Y}_{i}|\mathbf{A}_{i}, L_{i}) + \delta \cdot \frac{\mathbf{A}_{i} - \hat{\mathbb{E}}(\mathbf{A}_{i}|L_{i})}{\hat{f}(\hat{Q}_{\tau}(\mathbf{Y}_{i}|\mathbf{A}_{i}, L_{i})|\mathbf{A}_{i}, L_{i})}$$

Next, we calculate the estimator as

$$\frac{1}{n}\sum_{i=1}^{n}\frac{A_{i}-\hat{\mathbb{E}}(A_{i}|L_{i})}{\frac{1}{n}\sum_{i=1}^{n}(A_{i}-\hat{\mathbb{E}}(A_{i}|L_{i}))^{2}}\left[\widetilde{Q}_{\tau}(Y_{i}|A_{i},L_{i})-\hat{\mathbb{E}}(\widetilde{Q}_{\tau}(Y_{i}|A_{i},L_{i})|L_{i})\right].$$

## INFERENCE

- Inference is based on the efficient influence function after cross-fitting.
- We furthermore require the following terms to be  $o_p(n^{-1/2})$ :

$$\begin{split} & \mathbb{E}\left[\left(\hat{Q}_{\tau}(Y|A,\boldsymbol{L})-Q_{\tau}(Y|A,\boldsymbol{L})\right)^{2}\right],\\ & \mathbb{E}\left[\left(1-\frac{f(Q_{\tau}(Y|A,\boldsymbol{L})|A,\boldsymbol{L})}{\hat{f}(\hat{Q}_{\tau}(Y|A,\boldsymbol{L})|A,\boldsymbol{L})}\right)^{2}\right]^{1/2}\mathbb{E}\left[\left(\hat{Q}_{\tau}(Y|A,\boldsymbol{L})-Q_{\tau}(Y|A,\boldsymbol{L})\right)^{2}\right]^{1/2},\\ & \mathbb{E}\left[\left(\mathbb{E}(A|\boldsymbol{L})-\hat{\mathbb{E}}(A|\boldsymbol{L})\right)^{2}\right]^{1/2}\mathbb{E}\left[\left(\mathbb{E}(Q_{\tau}(Y|A,\boldsymbol{L})|\boldsymbol{L})-\hat{\mathbb{E}}(\hat{Q}_{\tau}(Y|A,\boldsymbol{L})|\boldsymbol{L})\right)^{2}\right]^{1/2},\\ & \mathbb{E}\left[\left(\mathbb{E}(A|\boldsymbol{L})-\hat{\mathbb{E}}(A|\boldsymbol{L})\right)^{2}\right] \quad (\text{if } \beta_{\tau}\neq 0). \end{split}$$

- Weaker than standard parametric assumptions, but still non-negligible.
- This is why our inferences are assumption-lean, rather than assumption-free

## SIMULATION STUDIES



## SIMULATION STUDIES

 $\blacksquare$  We considered inference for  $\beta_{\tau}$  in

$$Q_{ au}(Y^a|L) - Q_{ au}(Y^0|L) = eta_{ au} a$$
 for all  $a$ .

L is 4-dimensional multivariate normal.

2 settings:

- Binary exposure:  $\mathbb{P}(A = 1 | L) = \exp((-0.5 + 0.2L_1 0.4L_2 0.4L_3 + 0.2L_4))$ .
- Continuous exposure:  $A \sim \mathcal{N}(-0.5 + L_1 2L_2 2L_3 + L_4, 2^2)$ .

The outcome was generated according to

$$Y = 1 + A + \sin(L_1) + L_2^2 + L_3 + L_4 + L_3 \cdot L_4 + \epsilon,$$

where  $\epsilon \sim \text{Gamma}(k, \theta)$ .

- Nuisance parameters are estimated using 'grf', 'SuperLearner' and 'FKSUM' R-packages.
- We contrast the proposal with an oracle quantile regression and a naive plug-in estimator.

## SIMULATION STUDIES

Setting	estimator	au= 0.5					au= 0.9				
		bias	SD	SE	Cov		bias	SD	SE	Cov	
Bin.	Oracle	-0.0017	0.19	0.20	96.6	-	0.011	0.56	0.60	96.0	
	Plugin	-0.70	0.12	0.015	0.1		-0.64	0.22	0.036	1.6	
	TL-CF	0.012	0.22	0.25	97.2		0.14	0.68	0.63	91.4	
Cont.	Oracle	-0.0013	0.035	0.036	95.6	C	0.0010	0.10	0.11	94.6	
	Plugin	-0.17	0.064	0.016	0.5		-0.39	0.11	0.021	0.0	
	TL-CF	-0.011	0.044	0.042	92.9		0.012	0.14	0.10	85.3	

- Sample size n = 500, quantile  $\tau$ , 1000 simulations
- Oracle: correctly specified QR
- Plugin: Naive plug-in estimator
- TL-CF: Targeted Learning with 5-fold cross-fitting

- bias: Monte Carlo bias
- SD: Monte Carlo standard deviation
- SE: averaged estimated standard error
- Cov: coverage of 95% CI

## CONCLUSION



### CONCLUSION

- Assumption-lean modeling aims to make statistical / causal analyses more hygienic, by being clear about what we are estimating when the model is wrong.
- Obtain valid inferences, despite the use of flexible data-adaptive / machine learning algorithms, even when the model is wrong.
- By focusing on conditional quantiles, we can
  - tackle continuous exposures,
  - make better patient-specific treatment decisions, and
  - study treatment effect heterogeneity.

## REFERENCES (1)

Hines, O., Dukes, O., Diaz-Ordaz K., and Vansteelandt, S. (2021). Demystifying statistical learning based on efficient influence functions. The American Statistician, 1-48.

van der Laan, M. J., & Rose, S. (2011). Targeted learning: causal inference for observational and experimental data. Springer Science & Business Media.

Vansteelandt, S., & Dukes, O. (2022). Assumption-lean inference for generalised linear model parameters (with discussion). Journal of the Royal Statistical Society - B, 84, 657-685.

Vansteelandt, S. (2021). Statistical modeling in the age of data science. Observational Studies, 7, 217-228.

Vansteelandt, S., Van Lancker, K., Dukes, O. & Martinussen, T. Assumption-lean Cox regression. Journal of the American Statistical Association, in press.

## **REFERENCES** (2)



Baklicharov, G, Ley, C., Gorasso, V., Devleesschauwer, B., Vansteelandt, S. (2024). Assumption-Lean Quantile Regression. *arXiv preprint arXiv:2404.10495*.